Seminar:

Learning distribution-free anchored linear structural equation models in the presence of measurement error.

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Main contributions and outline

Main contributions

- Establish an identifiability of distribution-free anchored linear SEMs based on the geometry-faithfulness assumption.
- Propose a consistent algorithm to discover a latent structure in the presence of measurement error.
- Provide various numerical experiments and analysis of real galaxy data.

Introduction	Algorithm
Preliminaries	Numerical experiments
 Identifiability for distribution-free 	 Real data analysis
anchored linear SEMs	Conclusion

Introduction

Introduction

- Identifiability of directed acyclic graphical models (DAG) is usually achieved by posing additional assumptions. For example,
 - Causal minimality: True graph is a minimal structure that is Markov to its distribution.
 - ▶ Faithfulness: Conditional independence implies d-separation.
 - Distributional constraints: Gaussian errors with equal variance (Peters and Bühlmann, 2014), non-Gaussian errors (Shimizu et al., 2006), etc.
- The aforementioned identifiability results work under **causal sufficiency** regime, excluding the presence of latent variables.
- However, in many real-world setting, observed variables are **imperfect measures** of corresponding true variables.

Motivating example



Figure 1: Anchored DAG between the latent variables X and the observed variables Z.

- Figure 1 visualizes the relationship between the latent variables *X* and the observed variables *Z*.
- One can observe that X₁ and X₃ are d-separated(blocked) by X₂, while the statement becomes false if we replace X to Z.

Previous research

	Latent structure	Contamination	Distributional constraints
Halpern et al., 2015	$P(X) = \prod_{j} P(X_{j} \mid X_{\operatorname{Pa}(j)})$	$Z_j = f_j(X_j)$	X,Z are binary
Zhang et al., 2017	$X = BX + \epsilon$	$Z = X + \Psi$	ϵ, Ψ are Gaussian
Zhang et al., 2018	$X = BX + \epsilon$	$Z = X + \Psi$	ϵ is non-Gaussian
Saeed et al., 2020	$X = BX + \epsilon$	$Z_j = f_j(X_j)$	ϵ is Gaussian Known moment relationship
Liu et al., 2022	$P(X) = \prod_{j} P(X_{j} \mid X_{\operatorname{Pa}(j)})$	$Z_j = f_j(X_j)$	Likelihood is given
Ours	$X = BX + \epsilon$	$Z_j = f_j(X_j)$	Known moment relationship

Preliminaries

Directed acyclic graph

 \rightarrow (2) \rightarrow (3) \leftarrow (4) \rightarrow (5)

Figure 2: 5-node DAG example.

- A DAG G = (V, E) consists of a set of nodes V = {1, ..., p} and a set of directed edges E ⊂ V × V with no directed cycles. Its skeleton is an undirected graph obtained by removing directions in the edges.
- A set of **parents** of node *k*, denoted by Pa(*k*), consists of all nodes *j* such that (*j*, *k*) ∈ *E*.
- If there is a directed path *j* → · · · → *k*, then *k* is a **descendant** of *j*, and *j* is called an **ancestor** of *k*.
- A node k is a collider if there exists a triple (j, k, ℓ) such that j → k ← ℓ, and we say such triple generates a v-structure.

D-separation and d-connection

 $(1) \longrightarrow (2) \longrightarrow (3) \longleftarrow (4) \longrightarrow (5)$

Figure 3: 5-node DAG example.

- Two nodes *j* and *k* in DAG *G* are **d-connected by a node set** S ⊂ V if there exists a path *P* between *j* and *k* such that for every node *l* on the path *P*
 - 1. if ℓ is a collider, either ℓ or its descendant is in S,
 - 2. otherwise ℓ is not in S.
- If *j* and *k* are not d-connected by *S*, we say *j* and *k* are **d-separated by** *S*.
 - ▶ 1 and 5 are d-separated by \emptyset , {2}, {4}, {2, 4}.
 - ▶ 1 and 5 are d-connected by {3}.

Markov equivalence class



Figure 4: Markov equivalence class of \mathcal{G}_1 (\mathcal{G}_2 and \mathcal{G}_3) and its CPDAG.

- A Markov equivalence class (MEC) is a set of DAGs that encode the same set of d-separations.
- It is known that all DAGs in the MEC have the same skeleton and the same v-structures.
- A complete partially directed acyclic graph (CPDAG) is a unique representation of MEC.

Linear SEM

• The joint distribution generated by a **DAG model** (\mathcal{G}, P) can be factorized as follows:

$$P(X) = P(X_1, ..., X_p) = \prod_{j=1}^{p} P(X_j \mid X_{\mathsf{Pa}(j)}).$$
(1)

 A linear SEM is a special DAG model of (1) where the joint distribution of a linear SEM is defined by the following linear equations: For all *j* ∈ *V*,

$$X_{j} = \sum_{k \in \mathsf{Pa}_{(j)}} \beta_{kj} X_{k} + \epsilon_{j}, \tag{2}$$

where $(\epsilon_j)_{j \in V}$ are independent, but possibly not identical errors with mean 0 and variance $(\sigma_i^2)_{j \in V}$.

• The linear SEM in (2) can be restated as a matrix form:

$$X = BX + \epsilon. \tag{3}$$

We denote *L*(*G*, *B*, *F*) as the linear SEM in (3) where *B* is the edge weight matrix, *G* is the underlying true DAG, and *ε* ~ *F*.

Anchored linear SEM

- An anchored DAG model considers a DAG model with latent variables.
- In our framework, we consider an **anchored linear SEM**, special case of an anchored DAG model, as follows: For all *j* ∈ *V*,

$$Z_j = f_j(X_j)$$
 and $X \sim \mathcal{L}(\mathcal{G}, \mathcal{B}, (0, \Sigma_{\epsilon})),$ (4)

where $\Sigma_{\epsilon} = \text{diag}(\sigma_1^2, ..., \sigma_p^2)$, and $f_j : \mathbb{R} \to \mathbb{R}$ can be linear, non-linear, or even non-deterministic function.

- ▷ (Additive measurement error model) $f_j(X_j) = X_j + \psi_j$, where $\psi_j \sim (0, \eta_i^2)$.
- ▷ (Dropout model) $f_j(X_j) = X_j \psi_j$, where $\psi_j \sim \text{Bernoulli}(p)$.
- ▷ (Poisson transformation) $f_i(X_i) = \text{Poisson}(|X_i|)$.

Identifiability for distribution-free anchored linear SEMs

Geometry-faithfulness



Observe that X₁ and X₃ are d-separated by X₂ iff [Σ⁻¹]₁₃ = 0. In addition, X₂ and X₃ are d-separated by X₁ iff [Σ⁻¹]₂₃ = 0.

Geometry-faithfulness

Assumption 1. Geometry-faithfulness

Consider a linear SEM $\mathcal{L}(\mathcal{G}, B, (0, \Sigma_{\epsilon}))$ that generates P(X), i.e., $X \sim \mathcal{L}(\mathcal{G}, B, (0, \Sigma_{\epsilon}))$. Then, for any pair of nodes $j, k \in V$, and for any subset $S \subset V \setminus \{j, k\}$,

j and *k* are d-separated by *S* in $\mathcal{G} \iff \rho_{j,k,S} \propto [(\Sigma_{L,L})^{-1}]_{j,k} = 0$,

where $\Sigma = (I_p - B)^{-1} \Sigma_{\epsilon} (I_p - B)^{-\top}$, $L = S \cup \{j, k\}$, and $\rho_{j,k,S}$ is the partial correlation coefficient of X_j and X_k given X_S .

- Geometry-faithfulness ensures that partial correlations directly reflect d-separations and connections within the graph.
- Under the geometry-faithfulness assumption, j and k are d-separated by S if and only if the residuals obtained by projecting X_i and X_k onto X_S are orthogonal.

Identifiability for distribution-free anchored linear SEMs

Theorem 1. Identifiability for distribution-free anchored linear SEMs

Consider a distribution-free anchored linear SEM with $\mathcal{L}(G, B, (0, \Sigma_{\epsilon}))$. Then, model is identifiable up to the MEC if the followings are satisfied.

- (A1). The latent distribution P(X) is geometry-faithful to \mathcal{G} .
- (A2). The observed random variables satisfy the following condition: For all $j \in V$,

$$Z_j \perp\!\!\!\perp \{Z_1, ..., Z_p, X_1, ..., X_p\} \setminus \{Z_j, X_j\} \mid X_j.$$

(A3). For all $j, k \in V$, there exists a finite-dimensional vector δ_j of monomials in Z_j and a finite-dimensional vector δ_{jk} of monomials in Z_j and Z_k , such that their means can be mapped to the moments of the latent variables by continuously differentiable functions g_{jj} and g_{jk} , such that $\mathbb{E}[X_j] = g_j(\mathbb{E}[\delta_j]), \mathbb{E}[X_j^2] = g_{jj}(\mathbb{E}[\delta_{jj}]), \text{ and } \mathbb{E}[X_jX_k] = g_{jk}(\mathbb{E}[\delta_{jk}]), \text{ and their covariance satisfies <math>Cov(\delta_j, \delta_{jk}) < \infty$.

Algorithm

Algorithm 1. CPDAG learning algorithm for distribution-free anchored linear SEMs

Input: *n* i.i.d. observations from an anchored linear SEM, $Z^{1:n}$, transformation \mathcal{T} and *g* such that $Cov(X) = \mathcal{T}(\mathbb{E}[g(Z)])$, and significance level α .

Output: Complete Partial DAG (CPDAG), $\hat{\mathcal{G}}_{cp}$.

- 1. Estimate the mean of g(Z) from $Z^{1:n}$.
- 2. Estimate the covariance matrix $\hat{\Sigma}$ for latent variables using \mathcal{T} and g.
- 3. Estimate the partial correlations of X using $\hat{\Sigma}$.
- 4. Estimate a CPDAG using a constraint-based algorithm (e.g., the PC algorithm) by conducting a consistent partial correlation test with *α*.

Return: Estimated CPDAG, $\hat{\mathcal{G}}_{cp}$.

Under appropriate conditions, Algorithm 1 is guaranteed to be consistent.

Consistency

Assumption 2. Strong geometry-faithfulness

Consider an anchored linear SEM (4). For any $j, k \in V$ and $S \subset V \setminus \{j, k\}$, if j and k are d-separated by S, then the corresponding partial correlation is zero; otherwise, it is bounded below by a constant. Precisely, there exists a constant $\rho_{\min} > 0$ such that

 $0 < \rho_{\min} \leq \inf \{ |\rho_{j,k,S}| : j \text{ and } k \text{ are d-connected by } S \}.$

Theorem 2. Consistency

Consider an anchored linear SEM (4) with the true CPDAG \mathcal{G}_{cp} . Suppose that the strong geometry-faithfulness assumption specified in Assumption 2 and Assumptions (A2)-(A3) in Theorem 1 are satisfied. Additionally, suppose that the sequence of significance level { $\alpha_n : n \in \mathbb{N}$ } satisfies $2(1 - \Phi(0.5\rho_{\min}\sqrt{n-p-1})) < \alpha_n < 1$ for all $n \in \mathbb{N}$ where $\Phi(\cdot)$ denotes the CDF of N(0, 1). Then, Algorithm 1 is consistent, i.e.,

$$\hat{\mathcal{G}}_{cp} \to G_{cp}$$
 as $n \to \infty$,

where $\hat{\mathcal{G}}_{cp}$ is the estimated CPDAG by Algorithm 1.

Numerical experiments

Experiment setting

- The number of its parents was uniformly selected from the set {1, ..., d_{in}}.
- β_{jk} in Equation (2) were randomly chosen to fall within the intervals $(-0.75, -0.25) \cup (0.25, 0.75)$.
- Models were constructed with various noise distributions: (i) Gaussian, (ii) Uniform, (iii) Student's t, and (iv) Discrete uniform.
- For the contamination process, (i) dropout models, (ii) additive measurement error models were considered.
 - ▶ In dropout models, the dropout probability was set to $\gamma = 0.1$.
 - ▶ In additive measurement error models, the variance of the measurement error was set at $\eta^2 = 0.25$.
- Finally, the significance level was set at $\alpha_n = 1 \Phi(n^{1/5}/2)$, respecting the theoretical result from Theorem 2.

Dropout models



Additive measurement error models



Real data analysis

Data description



Figure 7: The true DAG and CPDAG of variables within galaxy data.

- The data (*p* = 6, *n* = 3462) comprises galaxy brightness measurements, which also includes measurement errors for each variable.
 - ▶ Rmag (0.0069), mumax (0), ApDRmag (0), Mcz (0.0038), BbMAG (1.5233) and rsMAG (1.6508).

Results



Figure 8: CPDAGs estimated by the proposed, PC, and GES algorithms.

- The proposed algorithm successfully detects all true edges, while falsely detects an undirected edge (BbMAG, rsMAG).
- Contrary to the proposed method, the PC and GES algorithms fail to recover the true edges.

Discussion

Discussion

- Introduce the geometry-faithfulness assumption.
- Present a consistent learning algorithm based on the PC-algorithm.

- Known moment relationships between observed and latent variables is required.
- Develop high-dimensional models and DAG recovery algorithms.

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